

On some classical properties of doubly truncated mixture of burr XII and weibull distributions

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Abstract: Limited work has been conducted on the doubly truncation for the mixture of different distributions. This paper is concerned with the doubly truncated mixture of Burr XII and Weibull distributions. In this paper, classical properties of the doubly truncated mixture of Burr XII and Weibull distributions have been derived. Cumulative distribution function, hazard rate, failure rate, inverse hazard function, odd function and the cumulative hazard function, rth moment about origin, moment generating function characteristic function, moments about origin and mean, mean and variance, measure of skewness and kurtosis have been discussed.

Keywords: Doubly Truncation, Mixture, Weibull and Burr XII Distributions, Characteristic Function

Introduction

Truncated distribution is the type of conditional distribution which is obtained by restricting the range of the other distribution. It arises in the practical situations where the ability to record the data or observation is limited to the values which lie above or below the given cut point.

The probability density function for the doubly truncated distribution is

$$f_{DT}(x) = \frac{f(x)}{\int_a^b f(x) dx}, \quad a < x < b \quad (1)$$

Where $f(x)$ is the probability density function of the distribution

Truncation results in the data which are restricted above or below that result in the truncated sample. The truncation is applied to the particular distribution which yields in the new distribution and not belonging to that class of distribution. (B.R et al, 2000) in his study showed that the truncated distributions are utilized in many practical situations, especially in different industrial arrangements. Fuzzy truncated normal distribution with the applications of real life time data were completely discussed by H.Agahi and M.Ghezelayagh (2009). The estimation techniques of

the maximum likelihood for the singly and doubly truncated normal distribution were firstly considered by Cohen, A. C (1950, 1991).

1. Probability Density Function

The probability density function for the doubly truncated mixture of Burr XII and Weibull distributions can be expressed in the following form.

$$f_{DT}(x) = \frac{p_1 c k s^{-c} x^{c-1} \left[1 + \left(\frac{x}{s} \right)^c \right]^{-k-1} + p_2 \frac{\alpha}{\beta^\alpha} x^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha}}{\Lambda} \quad (1.1)$$

Where

$$\Lambda = \int_a^b p_1 c k s^{-c} x^{c-1} \left[1 + \left(\frac{x}{s} \right)^c \right]^{-k-1} dx + \int_a^b p_2 \frac{\alpha}{\beta^\alpha} x^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha} dx$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

2. Cumulative Distribution Function

The cumulative distribution function for the doubly truncated mixture of burr XII and Weibull distributions is

defined as

$$F_{DT}(x) = \frac{\int_a^x f(x) dx}{\int_a^b f(x) dx} \quad (2.1)$$

The final expression for the CDF is given below

$$F_{DT}(x) = \frac{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]} \quad (2.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

3. Reliability Function

The reliability function is defined as follow

$$R_{DT}(x) = 1 - F_{DT}(x) \quad (3.1)$$

Putting equation (2.2) in (3.1), the required reliability function is defined as below

$$R_{DT}(x) = 1 - \frac{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]} \quad (3.2)$$

$$R_{DT}(x) = \frac{p_1 \left[\left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]} \quad (3.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

4. Hazard Function

Hazard function is defined as the ratio of the probability density function and reliability function and is expressed in the following form

$$h_{DT}(x) = \frac{f_{DT}(x)}{R_{DT}(x)} \quad (4.1)$$

The hazard function can be obtained by putting (1.1) and (3.2) in (4.1) and the final expression is mentioned below

$$h_{DT}(x) = \frac{p_1 c s^{-c} x^{-c-1} \left[1 + \left(\frac{x}{s} \right)^c \right]^{-k-1} + p_2 \frac{\alpha}{\beta} x^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha}}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]} \quad (4.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

5. Cumulative Hazard Function

The cumulative hazard function can be defined as

$$\wedge_{DT}(x) = -\log R_{DT}(x) \quad (5.1)$$

It is obtained by putting (3.2) in (5.1) and is expressed in the following form

$$\wedge_{DT}(x) = -\log \left[\frac{p_1 \left[\left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]} \right] \quad (5.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

6. Reversed Hazard Function

Reversed hazard rate can be defined as the ratio of the probability density function and the cumulative distribution function i.e

$$r_{DT}(x) = \frac{f_{DT}(x)}{F_{DT}(x)} \quad (6.1)$$

By putting (1.1) and (2.2) in (6.1) the expression obtained is as follow

$$r_{DT}(x) = \frac{p_1 c s^{-c} x^{-c-1} \left[1 + \left(\frac{x}{s} \right)^c \right]^{-k-1} + p_2 \frac{\alpha}{\beta} x^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha}}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]} \quad (6.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

7. Odds Function

The odds function denoted by $O(x)$ is the ratio of cumulative distribution function and the reliability function and has the following form:

$$O_{DT}(x) = \frac{F_{DT}(x)}{R_{DT}(x)} \quad (7.1)$$

The above expression can be obtained by putting (2.2) and (2.2) in (7.1) which as follow

$$O_{DT}^{(x)} = \frac{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{\left(\frac{x}{\beta} \right)^\alpha} \right\} \right]}{p_1 \left[\left\{ 1 + \left(\frac{x}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{\left(\frac{x}{\beta} \right)^\alpha} \right\} \right] - p_1 \left[\left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] - p_2 \left[\left\{ e^{\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]} \quad (7.2)$$

Where $a < x < b$, $k, s, c, \alpha, \beta > 0$, $p_1 + p_2 = 1$

8. Rth Moment about Origin

rth moment for the real valued function can be defined as

$$\mu_r' = E(x^r)$$

$$\mu_r' = \int x^r f_{DT}(x) dx$$

The final expression for the rth moment about origin is as follow

$$\mu_r'^e = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{p_1 k s^r p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{r}{c} + j + 1} + \left(\frac{a}{s} \right)^{\frac{r}{c} + j + 1} \right]}{\frac{r}{c} + j + 1} + \frac{p_2 \beta^r p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{r}{\alpha} + i + 1} - \left(\frac{a}{\beta} \right)^{\frac{r}{\alpha} + i + 1} \right]}{\frac{r}{\alpha} + i + 1} \right] \quad (8.2)$$

Where

$$I_1 = \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right]$$

and

$$I_2 = \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]$$

9. Raw Moments about Origin

Putting $r = 1, 2, 3$ and 4 in equation (8.1) the first four raw moments obtained are as follow

$$\mu_1'^e = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{1}{c} + j + 1} + \left(\frac{a}{s} \right)^{\frac{1}{c} + j + 1} \right]}{\frac{1}{c} + j + 1} + \frac{\beta p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{1}{\alpha} + i + 1} - \left(\frac{a}{\beta} \right)^{\frac{1}{\alpha} + i + 1} \right]}{\frac{1}{\alpha} + i + 1} \right] \quad (9.1)$$

$$\mu_2'^e = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s^2 p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{2}{c} + j + 1} + \left(\frac{a}{s} \right)^{\frac{2}{c} + j + 1} \right]}{\frac{2}{c} + j + 1} + \frac{\beta^2 p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{2}{\alpha} + i + 1} - \left(\frac{a}{\beta} \right)^{\frac{2}{\alpha} + i + 1} \right]}{\frac{2}{\alpha} + i + 1} \right] \quad (9.2)$$

$$\mu_3'^e = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s^3 p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{3}{c} + j + 1} + \left(\frac{a}{s} \right)^{\frac{3}{c} + j + 1} \right]}{\frac{3}{c} + j + 1} + \frac{\beta^3 p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{3}{\alpha} + i + 1} - \left(\frac{a}{\beta} \right)^{\frac{3}{\alpha} + i + 1} \right]}{\frac{3}{\alpha} + i + 1} \right] \quad (9.3)$$

$$\mu_4'^e = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s^4 p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{4}{c} + j + 1} + \left(\frac{a}{s} \right)^{\frac{4}{c} + j + 1} \right]}{\frac{4}{c} + j + 1} + \frac{\beta^4 p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{4}{\alpha} + i + 1} - \left(\frac{a}{\beta} \right)^{\frac{4}{\alpha} + i + 1} \right]}{\frac{4}{\alpha} + i + 1} \right] \quad (9.4)$$

Where

$$I_1 = \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right]$$

And

$$I_2 = \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]$$

10. Moments about Mean

$$\mu_1 = 0 \quad (10.1)$$

$$\mu_2 = \mu_2' - (\mu_1')^2 = \text{Variance} \quad (10.2)$$

$$\mu_2 = \frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s^2 p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{2}{c}+j+1} + \left(\frac{a}{s} \right)^{\frac{2}{c}+j+1} \right]}{\frac{2}{c}+j+1} + \frac{\beta^2 p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{2}{\alpha}+i+1} - \left(\frac{a}{\beta} \right)^{\frac{2}{\alpha}+i+1} \right]}{\frac{2}{\alpha}+i+1} \right] \left[\frac{k s p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{1}{c}+j+1} + \left(\frac{a}{s} \right)^{\frac{1}{c}+j+1} \right]}{\frac{1}{c}+j+1} + \frac{\beta p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{1}{\alpha}+i+1} - \left(\frac{a}{\beta} \right)^{\frac{1}{\alpha}+i+1} \right]}{\frac{1}{\alpha}+i+1} \right] \right]^2 \quad (10.3)$$

$$\left[\frac{1}{p_1 I_1 + p_2 I_2} \left[\frac{k s p_1 \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{1}{c}+j+1} + \left(\frac{a}{s} \right)^{\frac{1}{c}+j+1} \right]}{\frac{1}{c}+j+1} + \frac{\beta p_2 \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{1}{\alpha}+i+1} - \left(\frac{a}{\beta} \right)^{\frac{1}{\alpha}+i+1} \right]}{\frac{1}{\alpha}+i+1} \right] \right]^2 = \text{Variance}$$

Where

$$I_1 = \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right]$$

And

$$I_2 = \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]$$

The 3rd and 4th moments about mean can be calculated by putting μ_1' , μ_2' , μ_3' and μ_4' in the following equations respectively

$$\mu_3 = \mu_3' - 3\mu_2' \mu_1' + (\mu_1')^3$$

$$\mu_4 = \mu_4' - 4\mu_1' \mu_3' + 6(\mu_1')^2 (\mu_2') - 3(\mu_1')^4$$

11. Measure of Skewness and Kurtosis

Skewness is the measure of extent that the distribution leans to the one side of the mean and kurtosis is used to measure the flatness or peakedness of the probability curve.

Skewness and kurtosis are denoted by β_1 and β_2 respectively and is expressed in the following form

$$\beta_1 = \frac{(\mu_3')^2}{(\mu_2')^3} \quad (11.1)$$

And

$$\beta_2 = \frac{\mu_4}{(\mu_2)^2} \quad (11.2)$$

12. Moment Generating Function

The moment generating function can be defined as

$$M_X(t) = E(e^{tX})$$

Where

$$E(e^{tX}) = \int e^{tX} f_{DT}(x) dx \quad (12.1)$$

$$f_{DT}(x) = \frac{p_1 k s^{-c} x^{c-1} \left[1 + \left(\frac{x}{s} \right)^c \right]^{-k-1} + p_2 \frac{\alpha}{\beta^\alpha} x^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha}}{p_1 \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right] + p_2 \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]}$$

The final expression for the moment generating function is as follow

$$E(e^{tX}) = \frac{1}{p_1 I_1 + p_2 I_2} \sum_{i=0}^{\infty} \frac{(it)^i}{i!} \left[\frac{p_1 k s^i \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{i}{c}+j+1} + \left(\frac{a}{s} \right)^{\frac{i}{c}+j+1} \right]}{\frac{i}{c}+j+1} + \frac{p_2 \beta^i \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{i}{\alpha}+m+1} - \left(\frac{a}{\beta} \right)^{\frac{i}{\alpha}+m+1} \right]}{\frac{i}{\alpha}+m+1} \right] \right] \quad (12.2)$$

Where

$$I_1 = \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right]$$

And

$$I_2 = \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]$$

13. Characteristic Function

Characteristic function is obtained by replacing t in moment generating function with it and is expressed as below

$$\varphi_X(t) = E(e^{itX}) \quad (13.1)$$

Its expression can be obtained by replacing t with it in the expression of moment generating function (12.2) which is

$$\varphi_X(t) = \frac{1}{p_1 I_1 + p_2 I_2} \sum_{i=0}^{\infty} \frac{(it)^i}{i!} \left[\frac{p_1 k s^i \sum_{j=0}^{\infty} (-1)^j \binom{k+1}{j} \left[\left(\frac{b}{s} \right)^{\frac{i}{c}+j+1} + \left(\frac{a}{s} \right)^{\frac{i}{c}+j+1} \right]}{\frac{i}{c}+j+1} + \frac{p_2 \beta^i \sum_{i=0}^{\infty} (-1)^i \left[\left(\frac{b}{\beta} \right)^{\frac{i}{\alpha}+m+1} - \left(\frac{a}{\beta} \right)^{\frac{i}{\alpha}+m+1} \right]}{\frac{i}{\alpha}+m+1} \right] \right] \quad (13.2)$$

Where

$$I_1 = \left[\left\{ 1 + \left(\frac{a}{s} \right)^c \right\}^{-k} - \left\{ 1 + \left(\frac{b}{s} \right)^c \right\}^{-k} \right]$$

And

$$I_2 = \left[\left\{ e^{-\left(\frac{a}{\beta} \right)^\alpha} \right\} - \left\{ e^{-\left(\frac{b}{\beta} \right)^\alpha} \right\} \right]$$

References

- [1] B.R. Cho, M.S. Govindaluri. (2002), Optimal Screening Limits in Multi-Stage Assemblies, International Journal Production Research 40, pp. 1993-2009.
- [2] H.Agahi and M.Ghezelayagh. (2009), Fuzzy Truncated Normal Distribution with Applications, International Journal of Applied Mathematics and Computation, 1(3), pp170-181.
- [3] Cohen, A. C. (1950), Estimating the mean and variance of normal populations from singly truncated and doubly truncated samples, Ann. Math. Statist., 21, pp 557-569.
- [4] Cohen, A. C. (1991), Truncated and Censored Samples, Marcel Dekker, New York